Post-processing of multi-model ensemble river discharge forecasts using censored EMOS

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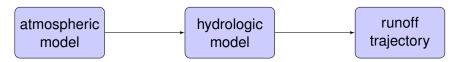
in cooperation with and funded by: Bundesanstalt für Gewässerkunde (BfG, Koblenz)²

EGU: 30 April 2014

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Deterministic hydrologic forecasts

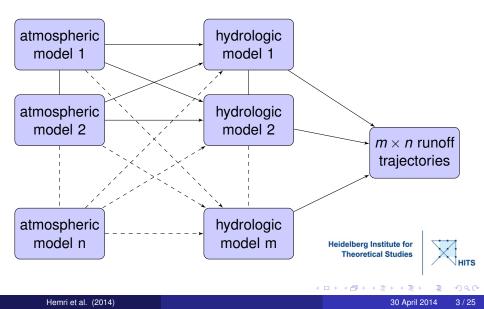


- Problem: statements on uncertainty impossible
- Generate ensemble forecasts by using several models, model configurations, and initial and boundary conditions.



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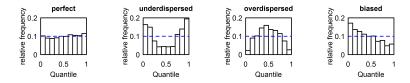
Hydrologic ensemble forecasts



Post-processing of ensemble forecasts

Main goals by Gneiting & Raftery (2007):

Well calibrated,



Hypothetical PIT histograms

and sharp probabilistic forecasts.

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- Standard post-processing methods rely mostly on continuously distributed variables (very often Gaussian).
- There is a need for a post-processing method for censored data:



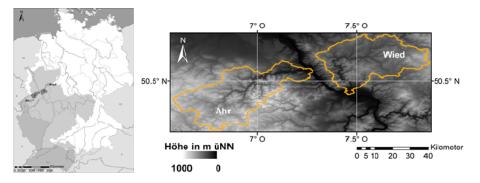
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- Standard post-processing methods rely mostly on continuously distributed variables (very often Gaussian).
- There is a need for a post-processing method for censored data: → Ensemble Model Output Statistics (EMOS) by Gneiting et al. (2005) is very suitable for this purpose due to its flexibility and simplicity

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Considered catchments

For testing the censored EMOS method we have selected the rivers Wied and Ahr:



Source: Bundesanstalt für Gewässerkunde (BfG), Koblenz (2013).

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Percentage of censored observations (obtained from the climatology from 1.11.1998 to 31.10.2008):

gauge	catchment	area [km²]	% censored
Altenahr (ALTE)	Ahr	746	72%
Friedrichsthal (FRIE)	Wied	680	55%



The hydrologic raw ensemble is obtained by running the HBV-96 model several times using the following meteorological input ensembles¹:

name	# members	lead-times	spatial resolution \sim
COSMO-LEPS	16	1-114 h	10 km
DWD-GME	1 (det.)	1-174 h	20 km
DWD-MER	1 (det.)	78 h (174 h)	7 km (20 km)
ECMWF-IFS	1 (det.)	1-240h	16 km

¹ DWD-MER stands for a model run based on COSMO-EU forcing up to lead-time 78 h and on DWD-GME thereafter.





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 \rightarrow A hydrologic raw ensemble of size 19 covering lead-times 1-114 h.

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Post-processing design

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- Hydrologic re-forecasts from 01.11.2008 to 25.10.2011.
- Post-process the forecasts for each lead-time separately.
- Seasonal training and verification periods:

verification period	training period
November 2008	SON 2009, SON 2010, SO 2011
DJF 2008/2009	DJF 2009/2010, DJF 2010/2011
MAM 2009	MAM 2010, MAM 2011

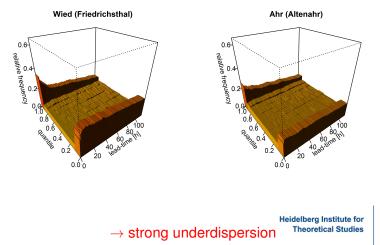
November 2008, SON 2009, SON 2010

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SO 2011

Calibration raw ensembles: 3D PIT





Hemri et al. (2014)

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Procedure:

- Box-Cox transformation in order to meet normality assumption
- subtract the censoring threshold
- use the following left-censored (Gneiting et al., 2004) and right-truncated distribution (Thorarinsdottir & Gneiting, 2010):

$$P(Y \le y \mid \overline{f}_1, \dots, \overline{f}_K) = \begin{cases} 0 & \text{if } y < 0\\ \frac{\Phi(\frac{y-\mu}{\sigma})}{\Phi(\frac{v-\mu}{\sigma})} & \text{if } 0 \le y \le v \\ 1 & \text{if } y > v \end{cases} \text{ where } \rightarrow$$

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- y: forecast runoff
- **\bar{f}_k:** model forecasts, mean value if from an ensemble
- v: upper threshold: 2 times the maximum of the climatology, also Box-Cox transformed → prevents unrealistic high forecast quantiles
- $\sigma^2 = c_1 + c_2 S^2$ (S^2 : variance among all raw forecast members) Heidelberg Institute for
- censoring threshold is set to zero
- $\mu
 ightarrow$ see next slide

Theoretical Studies

Due to the right-truncation the location parameter μ has to fulfill:

$$\mathbb{E}[Y|Y \le v] := \sum_{k=1}^{K} w_k \overline{f}_k + a \overline{\mathbb{1}_{f=0}} = \mu - \sigma \frac{\varphi(\frac{v-\mu}{\sigma})}{\Phi(\frac{v-\mu}{\sigma})}, \quad \text{where}$$

- *w_k*: weight of model *k*
- $\overline{\mathbb{1}_{f=0}}$: proportion of ensemble means \overline{f}_k that equal the lower threshold value (see Scheuerer, 2013)

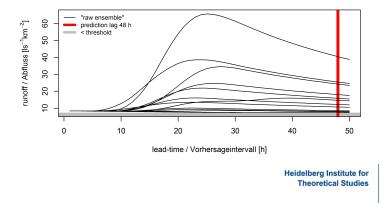
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$$\sigma = \sigma_0$$
 if $\mathbb{1}_{f=1}$

Parameters are estimated using minimal Continuous Ranked Probability Score (CRPS) estimation.



Censored EMOS: illustration I

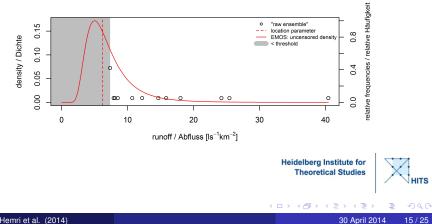
Example forecast initialized on 5.9.2009 06:00 CET:



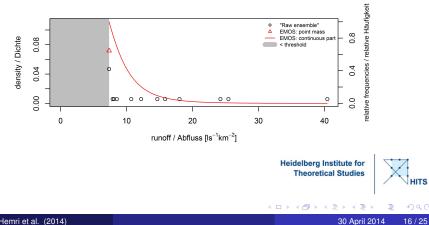
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Uncensored model pdf:

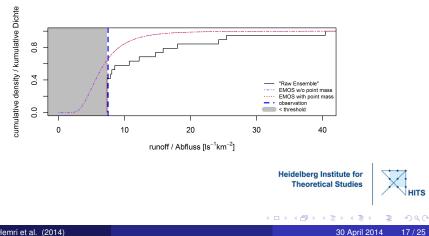


Censored model pdf:

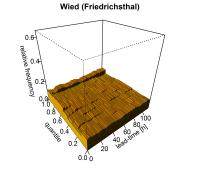


Censored EMOS: illustration IV

Censored model cdf:



Results: 3D PIT EMOS



Ahr (Altenahr)

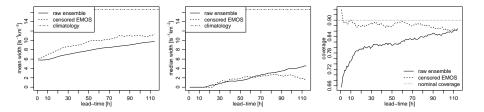
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Hemri et al. (2014)

From left to right: Mean prediction width, median prediction width and associated coverage of 90% prediction intervals:



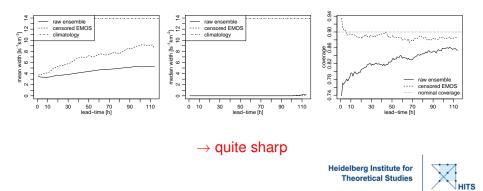
 \rightarrow quite sharp

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From left to right: Mean prediction width, median prediction width and associated coverage of 90% prediction intervals:



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Consider temporal dependencies among lead-times by:

- using a moving average of EMOS parameters (here: sliding window of size 5)
- using Copula approaches to consider correlation structure among lead-times:

→ Random Ensemble Copula Coupling that conserves the rank order structure of the raw ensemble (Schefzik et al., 2013) → Gaussian Copula approach (GCA) that conserves the

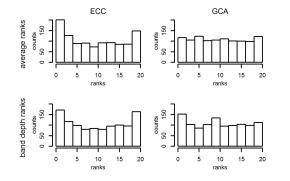
covariance structure estimated from the observations in the training period (Pinson & Girard, 2012)

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Results: multivariate calibration Wied

Average and band depth rank histograms (Thorarinsdottir et al., 2013):



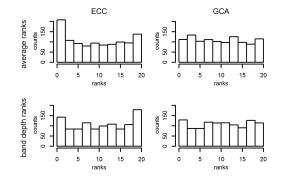
Heideberg Institute for GCA outperforms ECC in terms of correlation structure



Hemri et al. (2014)

Results: multivariate calibration Ahr

Average and band depth rank histograms (Thorarinsdottir et al., 2013):



Heidelberg Institute for GCA outperforms ECC in terms of correlation structure



Hemri et al. (2014)

- Statistical post-processing based on censored multi-model ensemble runoff forecasts yields appropriate predictive distributions.
- Censored EMOS improves calibration while not deteriorating sharpness much for the two examples considered.
- There exist straightforward methods for modeling of the temporal dependencies.
- GCA outperforms ECC in our example: → Are thus training observations better predictors of correlation structure than the raw ensemble?





- [1] T. Gneiting et al., Report to the Washington Technology Center, May 2004.
- [2] T. Gneiting et al., Monthly Weather Review, 133: 1098–1118, 2005.
- [3] T. Gneiting & A. E. Raftery, J. Am. Statist. Ass., 102: 359–378, 2007.
- [4] S. Hemri et al., *HyWa*, 58: 84–94, 2014.
- [5] P. Pinson & R. Girard, *Applied Energy*, 96 : 12–20, 2012.
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- [8] T. L. Thorarinsdottir & T. Gneiting, *Journal of the Royal Statistical Society* (*Series A*), 173 : 371–388, 2010.
- [9] T. L. Thorarinsdottir et al., *arXiv:1310.0236*, 2013.



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