Variations in the temporal evolution of seismicity pointed out by non-extensive statistical physics approach

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Statistics and pattern recognition applied to the spatio-temporal properties of seismicity

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Aim: to study a 'descriptor' of the criticality of a seismogenic system

the Earth's crust behaves as a **complex, nonlinear dynamical system** chracterized by **long-range** correlations and **multifractal** and hierarchical properties

 $\underline{similar}$ composite systems are studied in statistical physics through the Boltzmann-Gibbs (BG) entropy which is appropriate to study systems with \underline{local} interactions in space or time

Tool: we use the **Tsallis entropy** that generalizes some properties of BG entropy such as additivity

Newness:

- detailed analysis of the entropic *q* index and of the entropy for both **sub-additive** and **super-additive** states
- Bayesian estimation of the magnitude distribution in both cases

Brief notes on non-extensive statistical physics¹

Tsallis entropy of a probability density function f(x) is given by

$$S_q[f] = \frac{1 - \int f^q(x) \, dx}{q - 1},$$

where q is the entropic index.

 S_q is not additive, e.g. $S_q(A + B) = S_q(A) + S_q(B) + (1 - q)S_q(A)S_q(B)$ for any two independent systems; in particular, if q < 1 we speak of super-additivity and of sub-additivity when q > 1

We look for the distribution that optimizes the entropy S_q under the constraints

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$$\int_{0}^{+\infty} f(x) dx = 1$$
 normalization condition
• $\int_{0}^{+\infty} x f_q(x) dx = X_q$ where $f_q(x) = \frac{f^q(x)}{\int_{0}^{+\infty} f^q(x) dx}$ is called **escort** distribution,

maximizing the functional

$$\phi(f,\lambda_0,\lambda_1) = S_q + \lambda_0 \left(\int_0^{+\infty} f(x) \, dx - 1 \right) + \lambda_1 \left(\int_0^{+\infty} x \, f_q(x) \, dx - \bar{x} \right)$$

being λ_0 , λ_1 Lagrange multipliers.

¹more details on Bayesian inference in the last slide

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Q-exponential distributions

We obtain the q-exponential distributions

$$f(x) = rac{1}{X_q} \left[1 - rac{(1-q)}{(2-q)} rac{x}{X_q}
ight]^{1/(1-q)} \qquad 1 < q < 2 \qquad x \in \mathcal{R}^+$$

$$f(x) = \frac{1}{Z_q} \left[1 - (1 - q)\lambda' x \right]^{1/(1 - q)} \qquad 0 < q < 1 \qquad x \in (0, x_{max})$$

where Z_q is the normalizing constant, x_{max} is the cutoff to be imposed whenever the argument of $f(\cdot)$ becomes negative, $c_q = \int_0^{x_{max}} f^q(x) dx$ and $\lambda' = \frac{\lambda_1}{c_q + (1 - q) \lambda_1 \bar{x}}$

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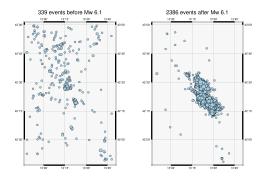
Q-exponential distributions for the magnitude

Following to the fragment-asperity interaction model^{*} [1], and applying the relationships $x \propto E^{2/3}$ and $\log_{10} M_0 = 1.5 M_w + 9.1$, we transform the previous q-exponential distributions into

$$\begin{split} \boxed{q \in (1,2)} & M \ge m_0 \\ & \frac{f(M)}{1 - F(m_0)} = \frac{\left(\ln 10\right) \frac{10^M}{a^{2/3} \bar{x}} \left[1 - \frac{1 - q}{2 - q} \frac{10^M}{a^{2/3} \bar{x}}\right]^{1/1 - q)}}{\left[1 - \frac{1 - q}{2 - q} \frac{10^{m_0}}{a^{2/3} \bar{x}}\right]^{(2 - q)/(1 - q)}} \\ \\ \boxed{q \in (0,1)} & M \in \left(m_0, \log_{10} \frac{a^{2/3}}{(1 - q) \lambda'}\right) \\ & \frac{f(M)}{1 - F(m_0)} = \frac{\left(\ln 10\right) (2 - q) \frac{\lambda'}{a^{2/3}} \times 10^M \left[1 - (1 - q) \frac{\lambda'}{a^{2/3}} 10^M\right]^{1/1 - q)}}{\left[1 - (1 - q) \frac{\lambda'}{a^{2/3}} 10^{m_0}\right]^{(2 - q)/(1 - q)}} \end{split}$$

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L'Aquila earthquake, M_w 6.1 on April 6, 2009 Lat (41.8, 43.0) Long (12.8, 13.8) 2005/04/25 - 2009/07/31



N = 2725 earthquakes $M_W \in (2.0, 6.1]$ drawn from the Italian Seismological Instrumental and Parametric Data-Base (ISIDe)

Figure 1. Spatial distribution of the seismicity, on the left, before and, on the right, after the main shock M_w 6.1

- We estimate the q-exponential distributions on time windows of 100 events shifted at each event of $\mathcal{M} = \left\{ M_w^{(i)} \right\}_{i=1}^N$
- we compare the two versions of the q-exponential distribution f(M) on the basis of the posterior marginal likelihood L(f(M))

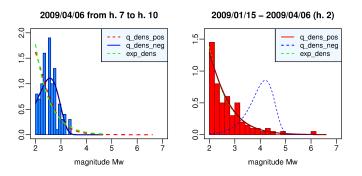


Figure 2. Time windows where (left) q < 1 and (right) q > 1 q-exponential distributions provide the "best" performance respectively

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Which is the **best** q-exponential density between q < 1 and q > 1 over all the time windows?

- in 96% of intervals log L(f_{q>1}(M)) > log L(f_{q<1}(M)), almost always showing strong evidence in favor of f_{q>1}(M) according to the Jeffreys' scale, that is, Δlog L(f(M)) > 2.3026
- in 70% of intervals $\log \mathcal{L}(f_{q>1}(\mathcal{M})) > \log \mathcal{L}(f_{q=1}(\mathcal{M}))$, of which 20% strongly and 46% substantially
- the 139 intervals in which $\log \mathcal{L}(f_{q<1}(\mathcal{M})) > \log \mathcal{L}(f_{q>1}(\mathcal{M}))$ are concentrated on the day of the mainshock

consequently, we show the values of the entropic index for the q-exponential distribution $f_{q>1}(M)$

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Temporal variations of the Q entropic index² (q > 1)

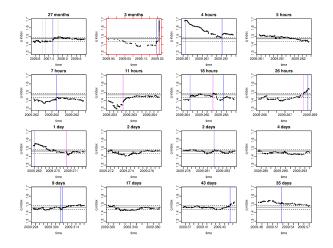


Figure 3. Vertical bars denote events of $M_w \in [4,5)$, $M_w \in [5,6)$, $M_w \ge 6$; horizontal, solid and dotted lines denote mean and quartiles of q values respectively.

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²in red the window covering the mainshock *Mw* 6.1

Temporal variations of the Tsallis entropy for the q > 1q-exponential distribution

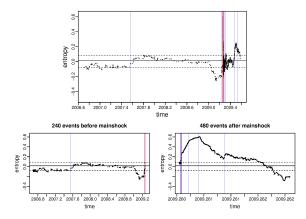


Figure 4. Solid and dotted lines denote mean and quartiles of q values respectively; (bottom) zoom around the date of L'Aquila shock.

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• about **q-index**

- background phase: q is about constant for a long period (2005 July 2007),
 - slight increase since M_w 4.1 on December 7, 2007 up to the end of 2008,
 - gradual decrease until March 28, 2009, then q increases monotonically since March 30, 2009 (M_w 4.1) up to q = 1.46 on the mainshock day
- aftershock sequence: q reachs the maximum ($q_{max} = 1.689$) within the 100 days following the mainshock
- caution must be paid to the different time scales: the longer the trend is, the more reliable the indication which it provides
- during the aftershock sequence the "average" decreasing trend is significant; looking at in detail the behavior - decrease pre-shock, increase after shock repeats during the sequence
- about the entropy
 - analogously, the entropy shows a slight increase after M_w 4.1 on December 7, 2007, it remains essentially constant for one year and a half, it starts to decrease since January 19, 2009 up to March 30, 2009, from here it increases (energy dispersion) rapidly up to the maximum 0.605 three hours after the mainshock
 - further significant increase after the pair $M_w 4.2 4.4$ within June 22, 2009 July 12. 2009 suggests that the sequence is in progress

Central Italy ... between L'Aquila and Amatrice (M_w 6 on 2016/08/24) earthquakes

Lat (42.3, 43.2) Long (12.7, 13.5)

2009/01/01 - 2016/10/30

N = 5673 earthquakes of $M_W \in (2.0, 6.5]$ drawn from the Italian Seismological Instrumental and Parametric Data-Base (ISIDe) considered as complete for $M_w \ge 2.0$

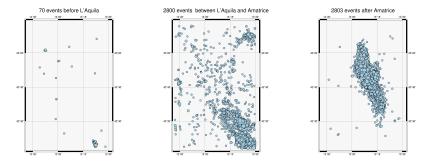


Figure 5. Spatial distribution of the seismicity, on the left, before L'Aquila shock, on the middle, between L'Aquila and Amatrice earthquakes, and, on the right, after Amatrice shock

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Temporal variations of the Q entropic index³ (q > 1)

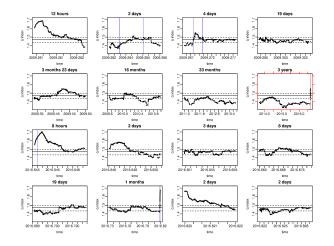
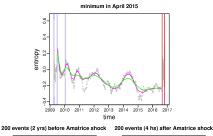


Figure 6. Vertical bars denote events of $M_w \in [4,5)$, $M_w \in [5,6)$, $M_w \ge 6$; horizontal, solid and dotted lines denote mean and quartiles of q values respectively.

³in red the window covering the mainshock Mw 6.0

Temporal variations of the Tsallis entropy for the q > 1q-exponential distribution



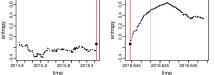


Figure 7. Smoothing of the entropy sequence, made through a locally weighted polynomial regression, highlights the **clear decrease** before Amatrice earthquake; (bottom) zoom around the date of Amatrice shock.

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• about **q-index**

- since the beginning of 2015, the value of q is constantly and significantly less than the mean $q_{mean} = 1.474$
- $q_{min} = 1.328$ on 2015/04/07

• about the entropy

• the global trend of the entropy is decreasing since the L'Aquila sequence, and a further decrease starts since August 2014

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Bayesian estimation: McMC methods

Let $\eta = (heta, eta)$ denote the parameter vector to estimate

in the Bayesian framework, quantities of interest (mean, variance, quantiles, ...) are obtained from the posterior distribution $p(\eta|\text{data})$ whose computation often involves intractable numerical integrations

to avoid these computational difficulties we resort to an iterative sampling Markov chain Monte Carlo method (McMC): the Metropolis-Hastings algorithm

 $\begin{array}{l} \label{eq:constraint} \hline \textbf{Metropolis-Hastings algorithm} \\ \bullet \mbox{ step } 1: \mbox{ select } \eta_0 \mbox{ from } \pi_0(\eta) \mbox{ and set } i = 1, \\ \bullet \mbox{ step } 2: \mbox{ draw a candidate } \tilde{\eta} \mbox{ from the proposal distribution } q(\eta \mid \eta_{i-1}), \\ \bullet \mbox{ step } 3: \mbox{ compute the acceptance probability} \\ & \alpha(\tilde{\eta} \mid \eta_{i-1}) = \min\left(1, \ \frac{\pi_0(\tilde{\eta}) \mbox{ } \mathcal{L}(\mbox{ data } \mid \tilde{\eta}) \mbox{ } q(\eta_{i-1} \mid \tilde{\eta})}{\pi_0(\eta_{i-1}) \mbox{ } \mathcal{L}(\mbox{ data } \mid \eta_{i-1}) \mbox{ } q(\tilde{\eta} \mid \eta_{i-1})}\right) \ , \\ \bullet \mbox{ step } 4: \mbox{ accept } \tilde{\eta} \mbox{ as } \eta_i \mbox{ with probability } \alpha(\tilde{\eta} \mid \eta_{i-1}), \mbox{ setp } \eta_i = \eta_{i-1} \mbox{ otherwise,} \\ \bullet \mbox{ step } 5: \mbox{ repeat steps } 2\text{-}4 \mbox{ a number R of times to get R draws from the posterior distribution, with optional burn-in and/or thinning.} \end{array}$

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